

WOULD YOU LIKE SOME PROBABILITY & STATISTICS?

$$\textcircled{1} e^x = 1 + x + \frac{x^2}{2!} + \frac{x^3}{3!} + \dots = \sum_{i=0}^{\infty} \frac{x^i}{i!}$$

$$\textcircled{2} a + ar + ar^2 + \dots = \sum_{i=0}^{\infty} ar^i = \frac{a}{1-r} \quad (\text{infinite sum})$$

$$\textcircled{3} a + ar + ar^2 + \dots + ar^n = \sum_{i=0}^n ar^i = a \left(\frac{1 - r^{n+1}}{1 - r} \right) \quad (\text{finite sum})$$

$$\textcircled{4} \int_{-\infty}^{\infty} f(x) dx = 1, \text{ where } f(x) \text{ is the density fnc. of a cnts distr.}$$

Geometric

$$X \sim \text{geometric}(p) \quad P(X=x) = p(1-p)^{x-1}$$

$$M_X(t) = \sum_{x=1}^{\infty} e^{tx} p(1-p)^{x-1} = pe^t \sum_{x=1}^{\infty} e^{t(x-1)} (1-p)^{x-1}$$

$$= pe^t \sum_{x=1}^{\infty} [(1-p)e^t]^{x-1} = pe^t \sum_{x=0}^{\infty} [(1-p)e^t]^x$$

$$\stackrel{(2)}{=} pe^t \left(\frac{1}{1-(1-p)e^t} \right) = \frac{pe^t}{pe^t - e^t + 1}$$

$$M'_X(0) = \frac{1}{p} = E(X)$$

$$M''_X(0) = \frac{2-p}{p^2}$$

$$\text{Var } X = \frac{1-p}{p^2}$$

Another way to calculate $E(X)$:

$$E(X) = \sum_{x=1}^{\infty} x p(1-p)^{x-1} = p \sum_{x=1}^{\infty} x (1-p)^{x-1}$$

$$= p \sum_{x=1}^{\infty} -\frac{d}{dp} [(1-p)^x] = -p \frac{d}{dp} \left(\sum_{x=1}^{\infty} (1-p)^x \right)$$

$$\stackrel{(2)}{=} -p \frac{d}{dp} \left(\frac{1-p}{1-(1-p)} \right) = -p \frac{d}{dp} \left(\frac{1}{p} - 1 \right) = -p \left(\frac{-1}{p^2} \right) = \frac{1}{p}$$

Gamma

$$X \sim \text{gamma}(\alpha, \lambda) \quad P(X=x) = \frac{\lambda^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-\lambda x}$$

$$M_X(t) = \int_0^\infty e^{tx} \frac{\lambda^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-\lambda x} dx = \int_0^\infty \frac{\lambda^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-(\lambda-t)x} dx$$

$\downarrow \quad \lambda^* = \lambda - t$

Let's make the inside of the integral look like the density fnc of gamma(α, λ^*)

density of gamma($\alpha, \lambda-t$)
 integrates to 1.
 \downarrow
 $\frac{(\lambda-t)^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-(\lambda-t)x}$

$$= \frac{\lambda^\alpha}{(\lambda-t)^\alpha} \int_0^\infty \frac{(\lambda-t)^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-(\lambda-t)x} dx \stackrel{(4)}{=} \left(\frac{\lambda}{\lambda-t} \right)^\alpha.$$

$$M'_X(0) = \frac{\alpha}{\lambda} = E(X)$$

$$M''_X(0) = \frac{\alpha(\alpha+1)}{\lambda^2}$$

$\left. \begin{array}{l} \\ \end{array} \right\} \text{Var } X = \frac{\alpha}{\lambda^2}$

Exponential

$$X \sim \text{exponential}(\lambda) = \text{gamma}(1, \lambda)$$

$$E(X) = \frac{1}{\lambda}$$

$$\text{Var } X = \frac{1}{\lambda^2}$$

$$\chi_n^2 \quad X \sim \chi_n^2 = \text{gamma}\left(\frac{n}{2}, \frac{1}{2}\right)$$

$$E(X) = n$$

$$\text{Var}(X) = 2n$$

Poisson

$$X \sim \text{Poisson}(s) \quad P(X=x) = \frac{e^{-s} s^x}{x!}$$

$$M_X(t) = \sum_{x=0}^{\infty} e^{tx} e^{-s} \frac{s^x}{x!} = e^{-s} \sum_{x=0}^{\infty} \frac{(se^t)^x}{x!}$$

$$\stackrel{(1)}{=} e^{-s} e^{set} = e^{s(et-1)}$$

$$\begin{aligned} M'_X(0) &= s = E(X) \\ M''_X(0) &= s+s^2 \end{aligned} \quad \left. \begin{aligned} \text{Var } X &= s \\ \text{interestingly, } E(X) &= \text{Var } X. \end{aligned} \right\}$$

Normal

$$X \sim \text{Normal}(\mu, \sigma^2)$$

$$M_X(t) = \int_{-\infty}^{\infty} e^{tx} \frac{1}{\sqrt{2\pi} \sigma} e^{\frac{-(x-\mu)^2}{2\sigma^2}} dx = \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi} \sigma} e^{\frac{-(x-\mu)^2 + tx}{2\sigma^2}} dx$$

$$\begin{aligned} \text{Notice that } -\frac{((x-\mu)-\sigma^2 t)^2}{2\sigma^2} &= \frac{-1}{2\sigma^2} \left((x-\mu)^2 - 2\sigma^2 t(x-\mu) + (\sigma^2 t)^2 \right) \\ &= \frac{-(x-\mu)^2}{2\sigma^2} + tx - \mu t - \frac{\sigma^2 t^2}{2}. \end{aligned}$$

$$\text{Hence, } \frac{-(x-\mu)^2}{2\sigma^2} + tx = \frac{-(x-\mu-\sigma^2 t)^2}{2\sigma^2} + \mu t + \frac{\sigma^2 t^2}{2}$$

$$M_X(t) = \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi} \sigma} e^{\frac{-(x-\mu-\sigma^2 t)^2}{2\sigma^2}} e^{\mu t + \frac{\sigma^2 t^2}{2}} dx$$

$$= e^{\mu t + \frac{\sigma^2 t^2}{2}} \int_{-\infty}^{\infty} \underbrace{\frac{1}{\sqrt{2\pi} \sigma} e^{\frac{-(x-\mu-\sigma^2 t)^2}{2\sigma^2}} dx}_{\text{density of Normal}(\mu+\sigma^2 t, \sigma^2)} \stackrel{(4)}{=} e^{\mu t + \frac{\sigma^2 t^2}{2}}$$

Bernoulli $X \sim \text{Bernoulli}(p)$

$$P(X=1) = p, \quad P(X=0) = 1-p.$$

$$M_X(t) = \sum_{x=0}^1 e^{tx} P(X=x) = e^{t \cdot 0} p(X=0) + e^{t \cdot 1} p(X=1) \\ = 1-p + pe^t.$$

$$\left. \begin{array}{l} M'_X(0) = p = E(X) \\ M''_X(0) = p \end{array} \right\} \text{Var } X = p(1-p).$$

Binomial $X \sim \text{Binomial}(n, p)$.

$X = X_1 + X_2 + \dots + X_n$, where each $X_i \sim \text{Bernoulli}(p)$ and indep.

$$E(X) = E\left(\sum_{i=1}^n X_i\right) = \sum_{i=1}^n E(X_i) = \sum_{i=1}^n p = np$$

$$\text{Var } X = \text{Var}\left(\sum_{i=1}^n X_i\right) \xrightarrow[\text{independence}]{\text{due to}} \sum_{i=1}^n \text{Var } X_i = np(1-p).$$

Negative Binomial $X \sim \text{neg. bin.}(p, j)$.

$X = \sum_{i=1}^j X_i$ for each $X_i \sim \text{geometric}(p)$ and indep.

$$E(X) = \sum_{i=1}^j E(X_i) = \sum_{i=1}^j \frac{1}{p} = \frac{j}{p}$$

$$\text{Var } X = \sum_{i=1}^j \text{Var } X_i = \frac{j(1-p)}{p^2}$$

due to indep.